

STAT 821 HOMEWORK 4 SOLUTION

Question 1

Proof: $X_1, \dots, X_n \stackrel{iid}{\sim} U(\theta - \frac{1}{2}, \theta + \frac{1}{2})$.

Let $Y_i = X_i - \theta + \frac{1}{2}$, $i = 1, \dots, n$, then $Y_1, \dots, Y_n \stackrel{iid}{\sim} U(0, 1)$ and $Y_{(1)} = X_{(1)} - \theta + \frac{1}{2}$, $Y_{(n)} = X_{(n)} - \theta + \frac{1}{2}$.

It's known that $Y_{(1)} \sim \text{Beta}(1, n)$ and $Y_{(n)} \sim \text{Beta}(n, 1)$. Thus

$$E(Y_{(1)}) = \frac{1}{n+1} \quad E(Y_{(n)}) = \frac{n}{n+1}$$

and

$$E(X_{(n)} - X_{(1)}) = E(Y_{(n)} - Y_{(1)}) = \frac{n-1}{n+1}$$

$$\text{i.e. } E\left(X_{(n)} - X_{(1)} - \frac{n-1}{n+1}\right) = 0$$

Let

$$h(I) = X_{(n)} - X_{(1)} - \frac{n-1}{n+1}$$

, then the distribution of $X_{(n)} - X_{(1)}$ does not depend on θ .

$$E(h(I)) = 0 \quad \forall \theta$$

However, $P(h(I) = 0) \neq 1$, so $T(X) = (X_{(1)}, X_{(n)})$ is not complete.

Question 2

$f(x, y)$ has continuous partial derivatives of the first and second order on \mathbb{R}^2 .

$$\nabla f(x, y) = \begin{pmatrix} 2x \\ 2y \end{pmatrix} \quad H(f(x, y)) = \begin{vmatrix} 2 & 0 \\ 0 & 2 \end{vmatrix}$$

$\det(H(f(x, y))) = 4 > 0$ and the matrix is diagonal with positive diagonal elements. Thus $H(f(x, y))$ is positive definite and $f(x, y)$ is convex. At point $(1, 0)$, the support hyperplane $L(X)$ is

$$\begin{aligned} L(X) &= f(1, 0) + \nabla f(1, 0) \cdot \langle x, y \rangle - \langle 1, 0 \rangle \\ &= 1 + \langle 2, 0 \rangle \cdot \langle x - 1, y \rangle \\ &= 2x - 1 \end{aligned}$$

Question 3

Proof: First show $-l(w)$ is convex.

$$-l(w) = \sum_{i=1}^n \log[1 + \exp(-y_i w^T x_i)]$$

$$\nabla(-l(w)) = \sum_{i=1}^n \begin{bmatrix} \frac{(-y_i x_{i1}) \exp(-y_i w^T x_i)}{1 + \exp(-y_i w^T x_i)} \\ \vdots \\ \frac{(-y_i x_{ik}) \exp(-y_i w^T x_i)}{1 + \exp(-y_i w^T x_i)} \end{bmatrix}$$

$$H(-l(w)) = \sum_{i=1}^n \begin{bmatrix} \frac{(-y_i x_{i1})^2 \exp(-y_i w^T x_i)}{[1 + \exp(-y_i w^T x_i)]^2} & \cdots & \frac{(-y_i)^2 x_{i1} x_{ik} \exp(-y_i w^T x_i)}{[1 + \exp(-y_i w^T x_i)]^2} \\ \vdots & & \vdots \\ \frac{(-y_i)^2 x_{i1} x_{ik} \exp(-y_i w^T x_i)}{[1 + \exp(-y_i w^T x_i)]^2} & \cdots & \frac{(-y_i x_{ik})^2 \exp(-y_i w^T x_i)}{[1 + \exp(-y_i w^T x_i)]^2} \end{bmatrix}$$

$$= \sum_{i=1}^n \frac{\exp(-y_i w^T x_i)}{[1 + \exp(-y_i w^T x_i)]^2} x_i x_i^T$$

$H(-l(w))$ is positive definite since $x_i x_i^T$ is positive definite and hence $-l(w)$ is strictly convex.

Let

$$f(w_0) = -l(w_0) = -m$$

be the minimum and w_0 denote MLE. Suppose $\exists w_1$ s.t.

$$f(w_1) = f(w_0) = -m \quad \text{and} \quad w_1 \neq w_0$$

Then $\forall 0 \leq r \leq 1$,

$$rf(w_0) + (1-r)f(w_1) > f(rw_0 + (1-r)w_1)$$

by convexity of $-l(w)$. In other words,

$$-m > f(rw_0 + (1-r)w_1)$$

This is a contradiction to the fact that $-m$ is the minimum of $f(w)$.

Thus MLE is unique.

Question 4

Proof:

$$\begin{aligned} E_F[h(x)] &= \int h(x) dF \\ &= \int \int_0^{h(x)} dt dF \\ &= \int_0^{h(x)} \int_{\{x \in R^k: h(x) > t\}} dF dt \\ &= \int_0^\infty F(h(x) > t) dt \end{aligned}$$

Similarly

$$E_G[h(x)] = \int_0^\infty G(h(x) > t) dt$$

We have

$$\begin{aligned} &E_G[h(x)] - E_F[h(x)] \\ &= \int_0^\infty [G(h(x) > t) - F(h(x) > t)] dt \\ &= \int_0^\infty [1 - G(h(x) < t)] - [1 - F(h(x) < t)] dt \\ &= \int_0^\infty [F(h(x) < t) - G(h(x) < t)] dt > 0 \quad (*) \end{aligned}$$

Notice that the set $H = \{x : h(x) \leq t\}$ is a convex set. This is because $\forall x, y \in H$

$$h(rx + (1-r)y) \leq rh(x) + (1-r)h(y) \leq t$$

So $rx + (1-r)y \in H$.

Thus, (*) implies there is a convex set $A \in R^k$, with $\underline{0} \in A$ s.t. $\forall t_0 \in A$,

$$\begin{aligned} F(h(x) \leq t_0) - G(h(x) \leq t_0) &\geq 0 \\ \Rightarrow F(A) &\geq G(A) \quad \text{for such } A \end{aligned}$$

Question 5

(a)

$$\begin{aligned}\nabla f(x, y) &= \begin{pmatrix} -\alpha x^{\alpha-1} y^{1-\alpha} \\ -x^\alpha (1-\alpha) y^{-\alpha} \end{pmatrix} \\ H(f(x, y)) &= \begin{pmatrix} -\alpha(\alpha-1)x^{\alpha-2}y^{1-\alpha} & -\alpha(1-\alpha)x^{\alpha-1}y^{-\alpha} \\ -\alpha(1-\alpha)x^{\alpha-1}y^{-\alpha} & \alpha(1-\alpha)x^\alpha y^{-\alpha-1} \end{pmatrix}\end{aligned}$$

$$\det(H) = \alpha^2(1-\alpha)^2 x^{2\alpha-2} y^{-2\alpha} - \alpha^2(1-\alpha)^2 x^{2\alpha-2} y^{-2\alpha} = 0$$

$$\text{tr}(H) = \alpha(1-\alpha)x^{\alpha-2}y^{-\alpha-1}[x^2 + y^2] > 0$$

The eigenvalue λ 's are the roots of the equation

$$\lambda^2 - \lambda \text{tr}(H) + \det(H) = 0$$

$$\Rightarrow \lambda(\lambda - \text{tr}(H)) = 0$$

$$\Rightarrow \lambda = 0 \quad \text{or} \quad \lambda = \text{tr}(H) > 0$$

Both of the roots are non-negative and hence the Hessian matrix is positive semidefinite.

$$\Rightarrow f(x, y) = -x^\alpha y^{1-\alpha} \quad \forall 0 < \alpha < 1$$

is convex on $\{(x, y) \in \mathbb{R}^2 : x > 0, y > 0\}$.

(b) Use Jensen's Inequality

$$E(f(x)) \geq f(E(x))$$

for convex $f(\cdot)$. Here we have

$$f(x, y) = -x^\alpha y^{1-\alpha} \quad \forall 0 < \alpha < 1$$

So

$$E[-x^\alpha y^{1-\alpha}] \geq -(EX)^\alpha (EY)^{1-\alpha}$$

i.e.

$$E(X^\alpha Y^{1-\alpha}) \leq (EX)^\alpha (EY)^{1-\alpha}$$

(c) We first show that the natural parameter space

$$\Theta = \left\{ \eta : \int \exp(\eta' T) d\mu < \infty \right\}$$

is convex.

Suppose $\eta_1, \eta_2 \in \Theta$ and for $0 \leq r \leq 1$

$$\begin{aligned} & \int \exp(r\eta_1' T + (1-r)\eta_2' T) d\mu \\ \propto & E[(e^{\eta_1' T})^r (e^{\eta_2' T})^{1-r}] \\ \leq & [E(e^{\eta_1' T})]^r [E(e^{\eta_2' T})]^{1-r} \\ \propto & \left(\int \exp(\eta_1' T) d\mu \right)^r \left(\int \exp(\eta_2' T) d\mu \right)^{1-r} \\ < & \infty \quad \text{since } \eta_1, \eta_2 \in \Theta \end{aligned}$$

Therefore $r\eta_1 + (1-r)\eta_2 \in \Theta$ and hence Θ is convex.

$A(\eta)$ is defined on a convex set Θ since $Cov(T) = A(\eta)$ and $Cov(T)$ is positive semidefinite. Thus $A(\eta)$ is a convex function on Θ .